

**A Meta-Analysis of the Effect of**  
**Common Currencies on International Trade**

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## Data Set

- 34 studies estimate currency union effect on trade
- 754 point estimates of  $\gamma$

Estimates (of  $\gamma$  and standard error) taken from

$$\ln(\text{Trade}) = \gamma \text{CurrencyUnion} + \text{controls} + \text{error}$$

where CurrencyUnion a dummy (1 for countries in currency union)

## 34 Estimates of Effect of Currency Union on Trade

<b>Author</b>	<b>Year</b>	<b>g</b>	<b>s.e. of g</b>
<b>Rose</b>	2000	1.21	0.14
<b>Engel-Rose</b>	2002	1.21	0.37
<b>Frankel-Rose</b>	2002	1.36	0.18
<b>Rose-van Wincoop</b>	2001	0.91	0.18
<b>Glick-Rose</b>	2002	0.65	0.05
<b>Persson</b>	2001	0.506	0.257
<b>Rose</b>	2001	0.74	0.05
<b>Honohan</b>	2001	0.921	0.4
<b>Nitsch</b>	2002b	0.82	0.27
<b>Pakko and Wall</b>	2001	-0.38	0.529
<b>Walsh and Thom</b>	2002	0.098	0.2
<b>Melitz</b>	2001	0.7	0.23
<b>López-Córdova, Meissner</b>	2003	0.716	0.186
<b>Tenreyro</b>	2001	0.471	0.316
<b>Levy Yeyati</b>	2003	0.5	0.25
<b>Nitsch</b>	2002a	0.62	0.17

<b>Flandreau and Maurel</b>	2001	1.16	0.07
<b>Klein</b>	2002	0.50	0.27
<b>Estevadeoral, et al</b>	2003	0.293	0.145
<b>Alesina, Barro, Tenreyro</b>	2003	1.56	0.44
<b>Smith</b>	2002	0.38	0.1
<b>Bomberger</b>	2002	0.08	0.05
<b>Melitz</b>	2002	1.38	0.16
<b>Saiki</b>	2002	0.56	0.16
<b>Micco, Stein, Ordonez</b>	2003	0.089	0.025
<b>Kenen</b>	2002	1.222	0.305
<b>Bun and Klaassen</b>	2002	0.33	0.1
<b>de Souza</b>	2002	0.17	0.24
<b>de Sousa and Lochard</b>	2003	1.21	0.12
<b>Flam and Nordström</b>	2003	0.139	0.02
<b>Barr, Breedon and Miles</b>	2003	0.25	0.033
<b>de Nardis and Vicarelli</b>	2003	0.061	0.027
<b>Rose</b>	2004	1.12	0.12
<b>Subramanian-Wei</b>	2003	0.732	0.08

## Meta Analysis

- Set of quantitative techniques for evaluating and combining empirical results from different studies.
- Different point estimates (one per study) of given coefficient treated as individual observations
- Can use this vector of estimates to:
  - estimate underlying coefficient of interest
  - test hypothesis that coefficient is zero
  - link estimates to features of the underlying studies
- Each study weighted equally

## Test of Zero Effect

- Test null hypothesis  $\gamma=0$ , pooling 34 point estimates (and standard errors)
- Test due to Fisher (1932), uses p-values from 34 underlying  $\gamma$  estimates
- Under null hypothesis, p-values are independently and randomly drawn from a normal  $[0,1]$  distribution,  $-2 \text{ Sum}[\ln(p_i)]$  is chi-squared
- Test statistic: 1272  $\sim$  chi-squared(68) under  $H_0$ .
  - Clear rejection of null hypothesis of no effect!

## Meta-Estimate of $g$ Pooled across Different Studies

	<b>Pooled Estimate of <math>g</math></b>	<b>Lower Bound of 95% CI</b>	<b>Upper Bound of 95% CI</b>	<b>P-value for test of no effect</b>
<b>Fixed</b>	.29	.27	.31	.00
<b>Random</b>	.64	.51	.77	.00
<b>Fixed without Rose</b>	.22	.19	.24	.00
<b>Random without Rose</b>	.53	.40	.66	.00

**Table 1: Meta-Analysis of Currency Union Effect on Trade ( $g$ )**

## Findings

- Considerable heterogeneity
- Fixed and random effect estimators dissimilar
- *Economically big*; currency union increases trade > 25%
- No conclusions change if my six studies are dropped
  - Test-statistic rejects the hypothesis of no effect:  $721 \sim \text{chi-squared}(54)$   
under  $H_0$

## Influential Studies?

- o No single study is especially influential (studies omitted one by one):

Study Omitted:	Coefficient	95% CI, lower	95% CI, upper
Rose	.28	.26	.30
Engel-Rose	.29	.26	.31
Frankel-Rose	.28	.26	.30
Rose-van Wincoop	.28	.26	.31
Glick-Rose	.27	.25	.29
Persson	.29	.26	.31
Rose	.26	.24	.29
Honohan	.29	.26	.31
Nitsch	.29	.26	.31
Pakko-Wall	.29	.27	.31
Walsh-Thom	.29	.27	.31
Melitz	.29	.26	.31
Lopez-Cordova, Meissner	.29	.26	.31
Tenreyro	.29	.26	.31
Levy Yeyati	.29	.26	.31
Nitsch	.29	.26	.31

Flandreau and Maurel	.26	.24	.29
Klein	.29	.26	.31
Estevadeoral et al	.29	.27	.31
Alesina, Barro and Tenreyro	.29	.26	.31
Smith	.29	.26	.31
Bomberger	.30	.28	.32
Melitz	.28	.26	.30
Saiki	.29	.26	.31
Micco, Stein, Ordonez	.34	.31	.36
Kenen	.29	.26	.31
Bun and Klaassen	.29	.26	.31
de Souza	.29	.27	.31
de Sousa and Lochard	.28	.26	.30
Flam and Nordström	.35	.33	.38
Barr, Breedon and Miles	.29	.27	.32
de Nardis and Vicarelli	.30	.28	.33
Rose	.28	.26	.30
Subramanian-Wei	.28	.26	.30
Combined	.29	.27	.31

**Table 2: Sensitivity of Meta-Analysis of  $g$  to Individual Studies**



## Does Choice of “Preferred” Estimate Matter Much?

- Can use different estimates from (31) underlying studies (3 too small)
- All are economically large, economically significant

		<b>Pooled g Estimate</b>	<b>Lower 95% CI</b>	<b>Upper 95% CI</b>	<b>P-value of no effect</b>
<b>“Preferred”</b>	<b>Fixed</b>	.27	.25	.29	.00
<b>“Preferred”</b>	<b>Random</b>	.64	.51	.76	.00
<b>Median</b>	<b>Fixed</b>	.34	.31	.38	.00
<b>Median</b>	<b>Random</b>	.82	.62	1.01	.00
<b>25<sup>th</sup>-Percentile</b>	<b>Fixed</b>	.18	.15	.20	.00
<b>25<sup>th</sup>-Percentile</b>	<b>Random</b>	.52	.38	.67	.00
<b>10<sup>th</sup>-Percentile</b>	<b>Fixed</b>	.12	.10	.14	.00
<b>10<sup>th</sup>-Percentile</b>	<b>Random</b>	.37	.24	.51	.00
<b>5<sup>th</sup>-Percentile</b>	<b>Fixed</b>	.11	.10	.13	.00
<b>5<sup>th</sup>-Percentile</b>	<b>Random</b>	.38	.27	.49	.00

**Table 3: Sensitivity of Meta-Analysis of  $\gamma$  to Choice of “Preferred” Estimate**

## Which Study Characteristics drive Outcomes?

- Hard to do multivariate regression with 34 observations

<b>Study Characteristic</b>	<b>Slope Coefficient ( z-statistic )</b>	<b>Intercept ( z-statistic )</b>
Number of Observations in study	8.0 e-7 (.9)	.60 (6.7)
Focus on EMU Observations	-.55 (3.9)	.79 (10.6)
Short-Run Focus	-.42 (2.6)	.74 (9.3)
Standard Error of ?	.98 (1.4)	.49 (3.8)
Dummy for Rose as Author	.46 (2.8)	.54 (6.9)
Dummy for mainly cross-section or panel study	.46 (2.2)	.25 (1.2)
Number of Countries in study	.001 (1.6)	.46 (3.5)
Number of Years in study	.002 (0.4)	.59 (3.9)
Dummy for post-WWII study	-.10 (0.4)	.74 (3.0)

**Table 4a: Meta-Analysis: Bivariate Determination of g Across Studies**

Focus on EMU Observations	-.57 (2.8)	-.50 (4.0)
Dummy for Rose as Author	.40 (2.6)	.39 (2.8)
Dummy for mainly cross-section or panel study	.19 (1.0)	
Short-Run Focus	.14 (0.7)	
Intercept	.50 (2.9)	.69 (9.0)

**Table 4b: Meta-Analysis: Multivariate Determination of g Across Studies**

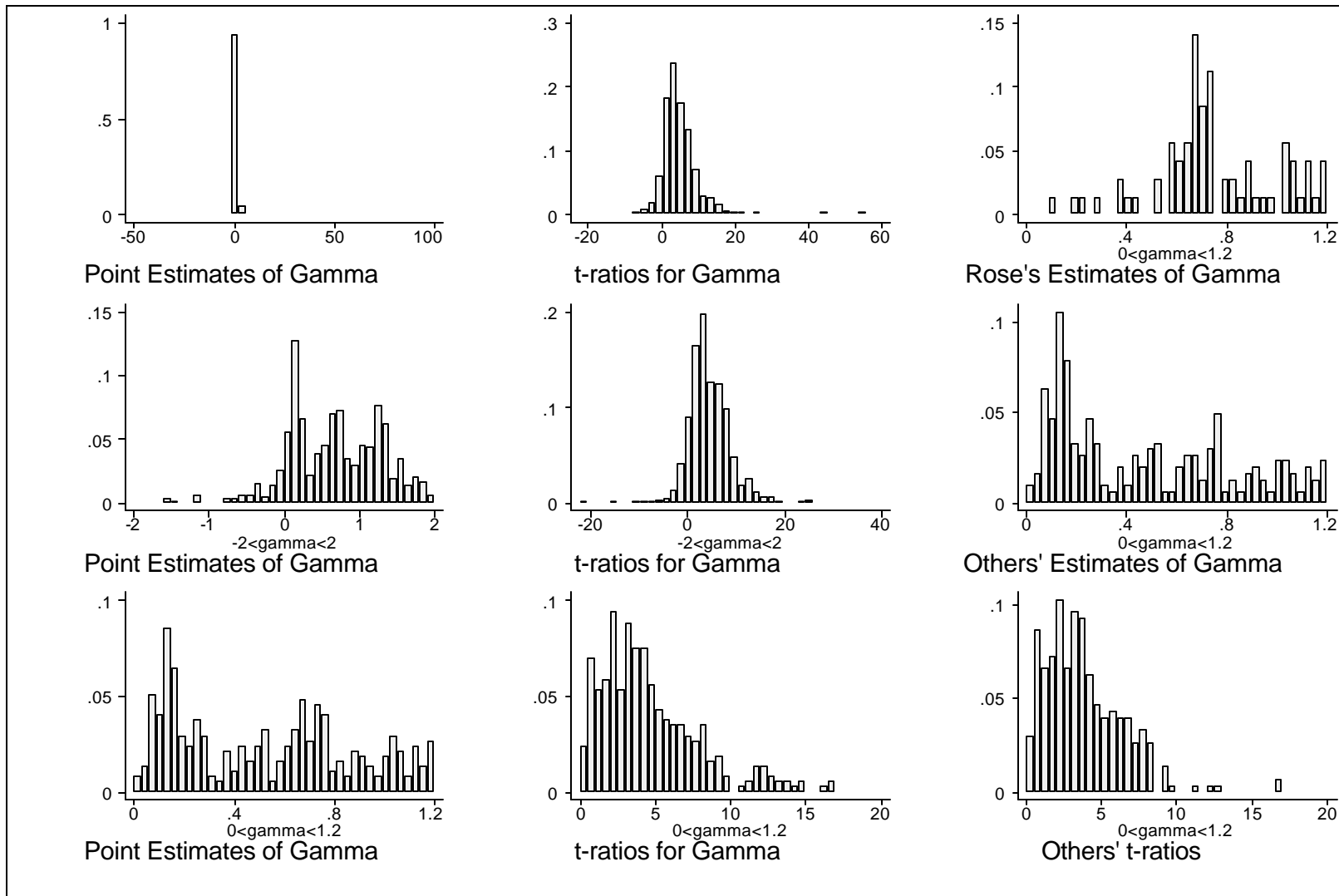
## Results

1. No positive relation between the number of observations and  $\gamma$ 
  - Worrying!
2. Papers that I co-author have higher point estimates
3. Euro studies have lower point estimates (too soon to tell? Smaller effect?)

## Different Estimates of $g$

- 34 studies but 754 estimates of  $\gamma$  (626 with standard errors)
  - Mean is .86
  - Mean t-ratio is 5.3
- Histograms of all estimates and t-statistics
- Vast majority are positive
  - Only 60 of the 754 (<8%) are negative
  - 325 (43%) exceed .69 (doubling)

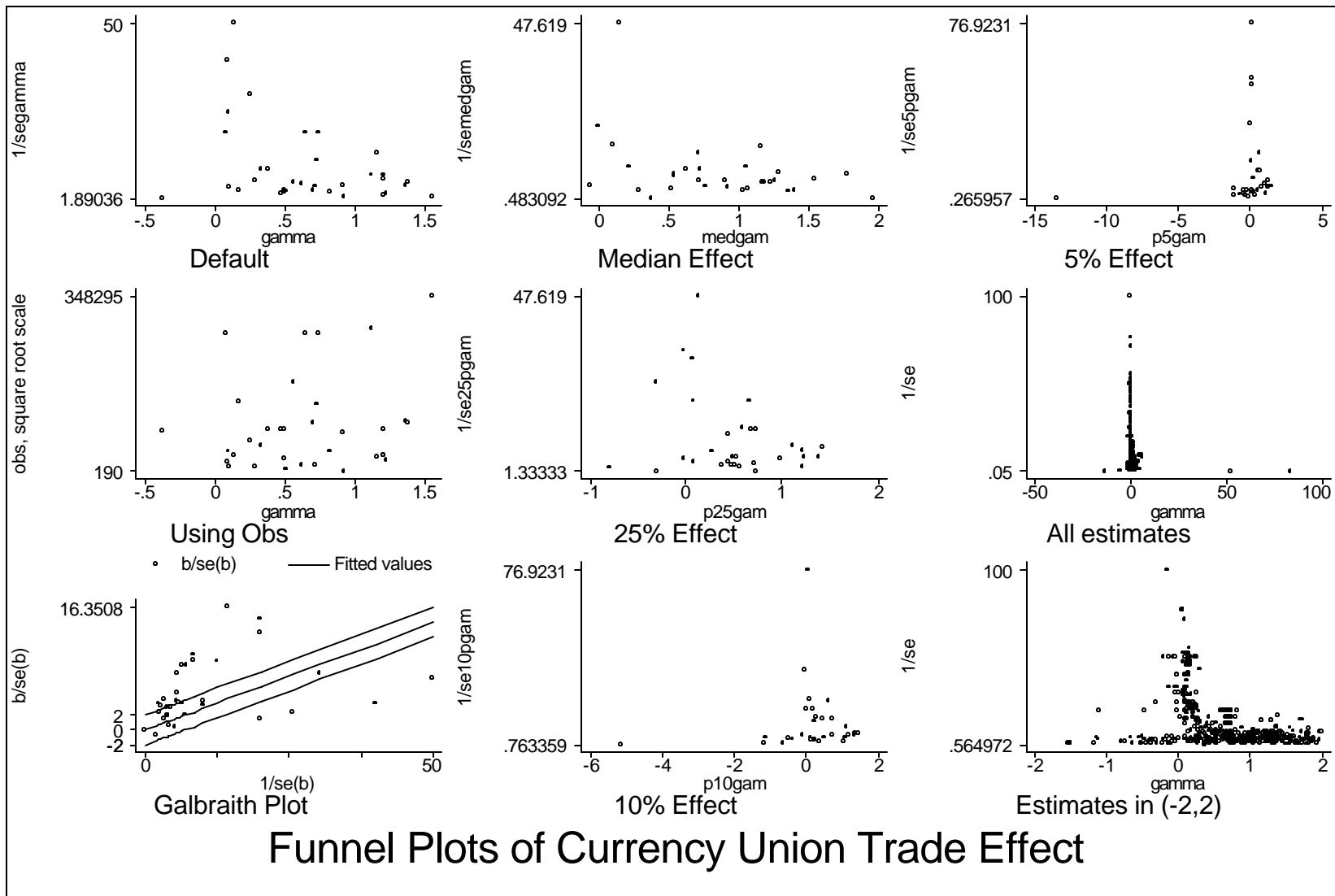
- Vast majority of  $\gamma$  estimates are statistically significant
  - Median t-statistic = 4.2
  - 479 (77%) exceed 2
  - Also true without Rose's estimates



**Figure 1: The Estimated Effect of Currency Union on Trade**

## Publication Bias

- “Funnel Plots” of  $\gamma$  against standard error seem asymmetric
  - Informal evidence of publication bias
  - “Galbraith Plot” of t-ratio for  $\gamma$  against standard error has many estimates outside confidence interval
  - Pervasive evidence of heterogeneity
- Begg and Mazumdar test ( $\gamma$  uncorrelated with its variance across studies) => publication bias
- Egger et al test: regress  $\gamma$  against precision (inverse s.e.) and bias also shows bias



**Figure 2: Funnel Plots for Publication Bias**



<b>Study</b>		<b>Coefficients</b>	<b>Coeff=0 (p-val.)</b>	<b>No. of Estimates</b>	<b>Heterogeneity (p-value)</b>
<b>Rose</b>	Fixed	1.289	0.000	52	0.00
	Random	1.311	0.000		
<b>Engel- Rose</b>	Fixed	1.350	0.000	5	0.78
	Random	1.350	0.000		
<b>Frankel-Rose</b>	Fixed	1.631	0.000	5	0.02
	Random	1.634	0.000		
<b>Rose-van Wincoop</b>	Fixed	0.230	0.000	18	0.00
	Random	0.649	0.000		
<b>Glick- Rose</b>	Fixed	0.697	0.000	37	0.00
	Random	0.772	0.000		
<b>Persson</b>	Fixed	0.647	0.000	6	0.11
	Random	0.586	0.000		
<b>Rose</b>	Fixed	0.824	0.000	17	0.00
	Random	1.060	0.000		
<b>Honohan</b>	Fixed	0.352	0.000	12	0.00
	Random	0.356	0.052		
<b>Nitsch</b>	Fixed	3.003	0.000	83	0.00
	Random	1.551	0.000		
<b>Pakko-Wall</b>	Fixed	0.874	0.000	6	0.00
	Random	0.332	0.350		
<b>Walsh-Thom</b>	Fixed	-0.008	0.574	7	0.00
	Random	0.020	0.542		
<b>Melitz</b>	Fixed	1.888	0.000	6	0.00
	Random	1.906	0.000		
<b>Lopez-Cordova and Meissner</b>	Fixed	0.723	0.000	47	0.38
	Random	0.722	0.000		
<b>Silvana Tenreyro</b>	Fixed	0.803	0.000	4	0.03
	Random	0.714	0.000		
<b>Levy Yeyati</b>	Fixed	1.014	0.000	19	0.02
	Random	1.055	0.000		
<b>Nitsch</b>	Fixed	0.464	0.000	8	0.00
	Random	0.429	0.009		
<b>Flandreau and Maurel</b>	Fixed	0.941	0.000	8	0.00
	Random	0.903	0.000		

<b>Klein</b>	Fixed	0.090	0.013	25	0.00
	Random	0.370	0.047		
<b>Estevadeoral, Frantz, and Taylor</b>	Fixed	0.433	0.000	18	0.01
	Random	0.450	0.000		
<b>Alesina, Barro and Tenreyro</b>	Fixed	1.159	0.000	8	0.00
	Random	1.649	0.000		
<b>Smith</b>	Fixed	1.007	0.000	17	0.00
	Random	1.118	0.000		
<b>Bomberger</b>	Fixed	0.205	0.000	6	0.00
	Random	0.315	0.006		
<b>Melitz</b>	Fixed	1.312	0.000	13	0.99
	Random	1.312	0.000		
<b>Saiki</b>	Fixed	1.162	0.000	16	0.00
	Random	0.520	0.008		
<b>Micco, Stein, Ordonez</b>	Fixed	0.098	0.000	54	0.00
	Random	0.130	0.000		
<b>Kenen</b>	Fixed	1.081	0.000	10	0.01
	Random	0.988	0.000		
<b>Bun and Klaassen</b>	Fixed	0.330	0.000	1	n/a
	Random	0.330	0.001		
<b>de Souza</b>	Fixed	-0.143	0.000	30	0.00
	Random	-0.018	0.714		
<b>de Sousa and Lochard</b>	Fixed	1.706	0.000	14	0.00
	Random	1.698	0.000		
<b>Flam and Nordström</b>	Fixed	0.150	0.000	49	0.00
	Random	0.149	0.000		
<b>Barr, Breedon and Miles</b>	Fixed	0.234	0.000	2	0.44
	Random	0.234	0.000		
<b>de Nardis and Vicarelli</b>	Fixed	0.090	0.000	2	0.90
	Random	0.090	0.001		
<b>Rose</b>	Fixed	0.905	0.000	10	0.00
	Random	0.988	0.000		
<b>Subramanian-Wei</b>	Fixed	1.142	0.000	11	1.0
	Random	1.142	0.000		

**Table 5: Within-Study meta-estimation of  $g$**

## Trade Diversion

- Does increased trade *inside* monetary unions divert trade away from non-members?

## Theory

- *Not analogous* to customs unions in welfare
- Trade diversion can be harmful because trade gains are less than lost tariff revenue
  - Ex: import goods at \$10, sell at \$15=\$10+\$5tariff
    - Can lose if eliminate tariffs from exporter with costs of \$12
  - But monetary union is simply a reduction in transactions costs; no lost tariff revenue (better bridges, not lower tolls)

## Practice

- Four Different Studies have searched for trade diversion
- *All* find evidence of *trade creation* between CU members & outsiders
  - Rose (2000): .29 (33%; se=.03)
  - Frankel and Rose (2002): .37 (45%; se=.04)
  - Micco et al: many EMU estimates in Tables 3-7: .086 (9%; se=.015)
  - Flam and Nordström: corroboration in Tables 3-7, .077 (8%; se=.017), growing over time and dis-aggregated too!

## Conclusion

- Too early to claim much
  - Studies are dependent and not all of equal interest
  - Estimates of  $\gamma$  are heterogeneous, cannot be linked to study features  
(especially precision/observations!)
- Publication Bias!
  - Intensely political issue (especially in Europe) => bias?
- Still, substantial evidence currency union has a positive effect on trade
- Effect is large economically, statistically
  - Currency union associated with trade effect in (30%, 90%)