

1 Why (18) is equivalent to (19)

1.1 Why (19) implies (18)

This direction is obvious.

1.2 Why (18) implies (19)

A little harder... First note that (19) implies that

$$\begin{aligned} E \left[\left\{ U'(x) + E[U'(x)]\sigma_m^{-2}R_m(\bar{R}_m - r) \right\} \sum_{j=1}^J b_{ji}\delta_j \right] &= 0. \\ \Rightarrow \sum_{j=1}^J b_{ji}E \left[\left\{ U'(x) + E[U'(x)]\sigma_m^{-2}R_m(\bar{R}_m - r) \right\} \delta_j \right] &= 0. \\ &\Rightarrow \sum_{j=1}^J b_{ji}c_j = 0. \end{aligned}$$

where

$$c_j = E \left[\left\{ U'(x) + E[U'(x)]\sigma_m^{-2}R_m(\bar{R}_m - r) \right\} \delta_j \right].$$

Thus (18) will hold so long as we can show that $c_i = 0 \forall i$. To show this assume the opposite, that some combination of the betas sum to 0 for every asset:

$$b_{ji} = \frac{\sum_{j=1}^J b_{ji}c_j}{c_j} \quad \forall i$$

Now using this equation, equation (3) in the paper can be written as

$$R_i = \bar{R}_i + \sum_{j=1}^{J-1} b_{ji}\delta_j - \delta_J \sum_{j=1}^J \frac{b_{ji}c_j}{c_j} + \epsilon_i.$$

Simplifying gives,

$$R_i = \bar{R}_i + \sum_{j=1}^{J-1} b_{ji} \left(\delta_j - \frac{c_j}{c_J} \delta_J \right) + \epsilon_i$$

$$R_i = \bar{R}_i + \sum_{j=1}^{J-1} b_{ji} \theta_j + \epsilon_i$$

where,

$$\theta_j = \left(\delta_j - \frac{c_j}{c_J} \delta_J \right)$$

Since the θ 's are also a factor structure,¹ this violates the assumption that the factor structure was minimal.

¹The θ 's do not satisfy the condition that they all have zero covariance with each other. However, this is no problem since they can be rotated to produce this result, i.e., the variance-covariance matrix is non singular so the new set of factors are $V^{-1}\theta$.