

How to divide the possession of a football?

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Abstract

The current National Football League overtime rule favors the team starting on offense. Auctioning off or dividing-and-choosing the starting possession can potentially restore *ex post* fairness. We find auctions to provide a better outcome when teams have asymmetric information.

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1. Introduction

In the National Football League (NFL), games ending in a tie are determined by sudden death overtime where the first team to score wins. Sudden death is an efficient means to decide a game which is violent and exhausting.² However, the sudden death nature confers a significant advantage on the team who has the first possession. While the outcome of a coin flip to determine first possession is *ex ante* fair, immediately after the toss it is no longer fair because the winning team has a significant chance of scoring on its first possession. NFL Commissioner Paul Tagliabue has said that “There has been a trend in the last seven or eight seasons that the team winning the toss in

overtime wins the game. That advantage of receiving the ball first is becoming unbalanced.” (New York Times, 2003a).³

Economic theory suggests natural solutions to this problem. One proposal would have the coaches of opposing teams bid on the yard line from its own end line, with the low bidder winning and starting its offense at the bid. Another proposal would have one team choose the starting yard line and the other team choose whether to cede or take possession at that yard line. We evaluate these proposals in terms of *ex post* fairness, both in the *complete information* setting in which teams have information about the chance of winning when starting at different yard lines as common knowledge and in *incomplete information* setting in which the information about the odds of winning associated starting at various yards is imperfect and dispersed between the two teams.

We find below that various recent proposals yield the same outcome that is *ex post* fair and envy-free in the complete information setting but that auctions provide a more *ex post* fair outcome than the divide-and-choose procedure when their information is incomplete.

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² When a college football game goes to overtime, each team is given one possession from its opponent’s twenty-five yard line. The leader after those possessions, if there is one, is declared the winner. If the teams remain tied, this continues, switching the order of possessions for each overtime, until one team leads the other at the end of the overtime.

³ Of the first 355 National Football League overtime games 100 were decided by the team winning the coin toss and then winning the game on its first possession (New York Times, 2003c). In 2002 of 25 overtime games, 10 were decided on the first possession (New York Times, 2003b) with 17 games eventually won by the team winning the coin flip.

2. Primitives

Suppose teams 1 and 2 tie in regulation, turning the game into an overtime. The continuation outcome at the beginning of overtime is characterized by the probability that team 1 would win. Let $a(x) + \alpha(s_1, s_2)$ and $b(y) + \beta(s_1, s_2)$ denote team 1's probability of winning when it begins offense at x from its end zone and when team 2 begins its offense at y from its end zone, respectively, where s_i is signal team $i=1, 2$ observes but its opponent does not. Obviously, $1 - b(y) - \beta(s_1, s_2)$ and $1 - a(x) - \alpha(s_1, s_2)$ will then represent team 2's probability of winning when it begins offense at y and when team 1 begins its offense at x , respectively. We assume that s_i is distributed over $[0, 1]$ according to a cdf $F_i(\cdot)$ which has a positive density f_i and that $\beta(\cdot, \cdot) - \alpha(\cdot, \cdot)$ is non-decreasing; i.e., an increase in the signals make each team prefer ceding a possession to having a possession at a given yard line. The signals thus capture the teams' private assessments about how the overtime game may unfold. Each team (and its coach) maximizes its winning probability.

It is reasonable to assume that $f(\cdot)$ is increasing and $g(\cdot)$ is decreasing; i.e., the further away from its end zone a team begins its offense, the higher is its chance to win. We also assume that a, b, α, β are continuous. Finally, we assume that $\forall (s_1, s_2), a(0) + \alpha(s_1, s_2) \leq b(0) + \beta(s_1, s_2)$ and $a(100) + \alpha(s_1, s_2) \geq b(100) + \beta(s_1, s_2)$. This last assumption means that a team would rather cede possession to the opponent team at its end zone than have a possession at its own end zone. Graphically, it simply means that the curves cross, as illustrated in Fig. 1.

Of special interest is the intersection $x^*(s_1, s_2)$ of the two curves. This value constitutes an envy-free solution in the sense that both teams are indifferent to a possession at that yardage. In fact, the problem of assigning an overtime possession is isomorphic to the well-known "cake-cutting" problem, except for the common value element in our problem.⁴ In the presence of incomplete information, an envy-free division may be difficult to implement, nor is it necessarily desirable. We rather focus on the following properties for a fair assignment rule.

- *Ex ante* symmetry: An assignment rule should treat both teams *ex ante* symmetrically in the sense that switching the teams in the way the rule applies should not affect the odds of winning for each team *ex ante*. This property is a minimum requirement of fairness. The current system of flipping coin is *ex ante* symmetric. So is any anonymous auction rule. The divide-and-choose rule, with a coin flip to determine the divider, is also *ex ante* symmetric.
- Non-arbitrariness: A fair assignment rule should allow the players' abilities and effort – rather than the way the rule assigns the possession – to determine the outcome. In particular, it should minimize the extent to which the fortune of each team is affected by the assignment. We say a rule is *non-arbitrary* if, for each state (s_1, s_2) , each team's winning

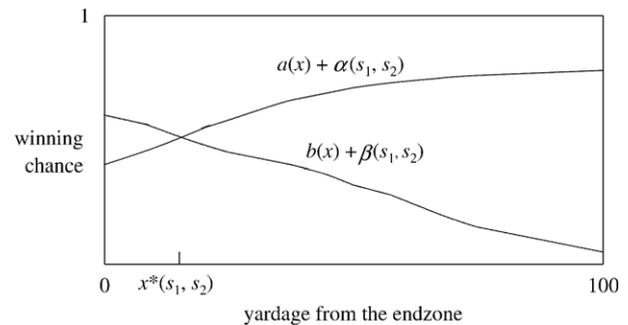


Fig. 1.

probability at the start of overtime play is non-random under that rule. Formally, a rule M is non-arbitrary if the highest winning probability $\bar{\phi}_M(s_1, s_2)$ and the lowest winning probability $\underline{\phi}_M(s_1, s_2)$ it induces for team 1 at the start of the play coincide for almost every (s_1, s_2) . This feature is precisely what is missing in the current system: The winning probability at the start of an overtime play is random, depending on the realized outcome of the coin flip. The proposed rules may or may not be non-arbitrary depending on the equilibrium play of the assignment games the rules induce.

Non-arbitrariness is certainly desirable. But it is unclear how to compare a non-arbitrary rule with an arbitrary rule, if they implement different winning probabilities on average. The desirability of the former will be clear-cut, however, if the winning probabilities implemented by the two rules are in the same range. We say *ex ante* symmetric and non-arbitrary rule A is *fairer* than an *ex ante* symmetric rule B if

$$\mathbb{E}_{(s_1, s_2)}[\underline{\phi}_B(s_1, s_2)] < \mathbb{E}_{(s_1, s_2)}[\phi_A(s_1, s_2)] < \mathbb{E}_{(s_1, s_2)}[\bar{\phi}_B(s_1, s_2)].$$

We are ready to analyze different rules. Below we consider two different types of rules: (a) γ -*auction*: Both teams bid simultaneously on the yardage, the lower bidder wins and has possession at $\gamma \cdot \text{low bid} + (1 - \gamma) \cdot \text{high bid}$, for some $\gamma \in [0, 1]$; (b) *Divide-and-choose*: A coin is flipped to choose a team; the team proposes a yardage; and the other team chooses between having a possession at that yard and giving a possession to other team at that yardage (from its own end line).

3. Benchmark: complete information

Suppose $\alpha(s_1, s_2) = \beta(s_1, s_2) \equiv 0$. In this case, the teams have completely common opinion about what the odds of winning will be for each team having possession at any given yard. Although this is very unrealistic, the complete information case delivers a very clean analysis and illustrates some desirable features of the recently proposed rules, so it can serve as a useful benchmark. We first characterize the equilibrium of the two proposed rules.⁵

⁴ In the standard cake cutting problems, each player is assumed to know the exact value of the cake, so the common value problem does not exist. See Steinhaus (1948).

⁵ Part of the analysis here was developed independently by Che (2003) and Kraska (2004).

Proposition 1. *In any γ -auction rule with $\gamma \in [0, 1]$, it is an equilibrium for each team to bid x^* such that $f(x^*) = g(x^*)$. The divide-and-choose rule has an equilibrium in which a team chosen (at random) to pick a yard proposes x^* . Team 1 wins with probability $f(x^*)$ in any equilibrium.*

Proof. Consider first a γ -auction, $\gamma \in [0, 1]$. Suppose team 2 bids x^* . Team 1's payoff from bidding any x is

$$\pi_1(x, x^*) = \begin{cases} a(\gamma x + (1 - \gamma)x^*) & \text{if } x < x^*, \\ \frac{a(x^*) + b(x^*)}{2} & \text{if } x = x^*, \\ b(\gamma x^* + (1 - \gamma)x) & \text{if } x > x^*. \end{cases}$$

Since $\frac{a(x^*) + b(x^*)}{2} = b(x^*) = a(x^*)$, $\pi_1(x, x^*)$ is non-decreasing x for any $x \leq x^*$ and non-increasing in x for $x > x^*$. So, the payoff is maximized at $x = x^*$. Hence, there is no profitable deviation for coach 1. A symmetric argument works for team 2, using the fact that $1 - a(x^*) = 1 - b(x^*)$.

Now consider the divide-and-choose rule. Suppose without loss team 1 is chosen to propose a starting yard line and team 2 decides on possession. If team 1 proposes x , team 2 will choose $\min\{1 - b(x), 1 - a(x)\}$, making team 1's payoff equal $\min\{a(x), b(x)\}$, graphed by the lower envelope of the two curves. Clearly, the payoff is maximized at x^* .

The uniqueness in the probability of winning follows from the fact that each game is a two-person zero sum game. \square

It follows that these rules are non-arbitrary. Under very weak conditions these rules are fairer than the current rule. In the current rule, the starting yard line for a team with a possession is a random variable resulting from the opponent's kickoff. It is a reasonable to assume.

Assumption A0. In the current system, a team that wins a coin toss chooses an offense; and there exists $x_0 \in X \cap Y$, $x_0 \neq x^*$, where X and Y are respectively the support of team 1 and team 2's starting position from one's own end line that results from receiving a kickoff.

This assumption is reasonable; in particular the second part simply assumes there to be some common starting position, differing from x^* , that arises with positive density from receiving a kickoff.

Corollary 1. *Given Assumption A0, a γ -auction and the divide-and-choose rule are fairer than the current system.*

Proof. Let ϕ_M is the probability of winning at the start of overtime play under $M=0$, γ , d representing respectively the current system, a γ -auction, and the divide-and-choose rule. Let $x_0 \in X \cap Y$, $x_0 \neq x^*$. Then,

$$\underline{\phi}_0 \leq \min\{a(x_0), b(x_0)\} < a(x^*) = \phi_\gamma = \phi_d < \max\{a(x_0), b(x_0)\} \leq \bar{\phi}_0.$$

Since the two rules are non-arbitrary, the proof is complete. \square

That these two rules implement the envy-free solution has an additional feature: *A better team has a higher chance to win.* To

see this, say team 1 is *more capable* than team 2 if $a(x) \geq 1 - b(x)$ for all x , and strictly so if the inequality holds strictly. That is, team 1 is more likely to win than team 2, if starting its offense at the same position or when facing an offense by its opponent team at the same yardage. Likewise, two teams are *equally capable* if $a(x) = 1 - b(x)$ for all x .

Corollary 2. *In each of the γ -auction and divide-and-choose rule, a (strictly) more capable team is (strictly) more likely to win, and equally capable teams have the equal chance of winning at the start of an overtime play in equilibrium.*

Proof. Let x^* be an equilibrium bid. Without any loss, suppose team 1 is more capable than team 2. Then, $a(x^*) \geq 1 - b(x^*)$. By Proposition 1, $a(x^*) = b(x^*)$. Combining the two facts, we have $a(x^*) \geq \frac{1}{2}$. The inequality is strict if team 1 is strictly more capable than team 2, and becomes an equality if two teams are equally capable. \square

Under the current system, two equally capable teams will not have the equal chance of winning after the assignment, except by a mere coincidence. Clearly, the equally capable teams have the same chance of winning prior to coin tossing, but what matters is the winning chance after the coin tossing, which is not necessarily a half even when the teams are equally capable.

4. Incomplete information

We now consider the general case where α and β need not vanish. We first characterize the equilibrium of the γ -auction.

Proposition 2. *Any equilibrium of γ -auction is a pure strategy profile $(x^\gamma(s_1), y^\gamma(s_2))$ that are strictly increasing.*

Proof. We first consider a γ -auction, and fix an equilibrium (possibly in mixed strategies). A standard argument shows that the support of the teams' equilibrium bids coincide and have no gaps and no mass points. Suppose team 2's equilibrium bid is distributed according to a cdf $G(y)$, and let $H(s_2|y)$ be the conditional distribution of type s_2 making bid y (i.e., $F_2(s_2) = \int_y H(s_2|y) dG(y)$). Since y has no mass point, team 1 with type s_1 receives payoff from bidding in the equilibrium support:

$$\begin{aligned} \Pi_1^\gamma(x; s_1) := & \int_{y \geq x} \int_{s_2} [a(\gamma x + (1 - \gamma)y) + \alpha(s_1, s_2)] dH(s_2|y) dG(y) \\ & + \int_{y < x} \int_{s_2} [b((1 - \gamma)x + \gamma y) + \beta(s_1, s_2)] dH(s_2|y) dG(y). \end{aligned}$$

Note that

$$\frac{\partial^2 \Pi_1^\gamma(x; s_1)}{\partial x \partial s_1} = \int_{s_2} \left[\frac{\partial(\beta(s_1, s_2) - \alpha(s_1, s_2))}{\partial s_1} \right] dH(s_2|x) > 0.$$

This means that the best response correspondence is monotonic (i.e., every selection is non-decreasing). It follows that the best response is a singleton for almost every s_1 . Since the same holds for team 2, we conclude that the equilibrium is in

pure strategies. No mass point implies also that the equilibrium strategies must be strictly increasing. \square

An immediate consequence of Proposition 2 is that a γ -auction is non-arbitrary. Again, a γ -auction is fairer than the current rule under a weak condition. Modify Assumption A0 by replacing X and Y with state specific $X(s_i, s_2)$ and $Y(s_i, s_2)$, and call it Assumption A0.

Corollary 3. *Given Assumption A0, a γ -auction is fairer than the current rule.*

Proof. Let $x_0 \in X(s_i, s_2) \cap Y(s_i, s_2), \forall (s_i, s_2)$. Let $(x^\gamma(\cdot), y^\gamma(\cdot))$ be an equilibrium of the γ -auction game. Since team 1 has an option of simply bidding x_0 in a γ -auction, $\forall (s_i, s_2)$,

$$\begin{aligned} \phi_\gamma(s_1, s_2) &\geq (a(\gamma x_0 + (1 - \gamma)y^\gamma(s_2)) + \alpha(s_1 + s_2)) \cdot 1_{\{x_0 \leq y^\gamma(s_2)\}} \\ &\quad + (b(\gamma y^\gamma(s_2) + (1 - \gamma)x_0) + \beta(s_1, s_2)) \cdot 1_{\{x_0 > y^\gamma(s_2)\}} \\ &\geq \min\{a(x_0) + \alpha(s_1, s_2), b(x_0) + \beta(s_1, s_2)\} \geq \underline{\phi}_0(s_1, s_2). \end{aligned}$$

The first inequality holds since x_0 need not be a best response and the second follows since $a(\cdot)$ is increasing and $b(\cdot)$ is decreasing. Since (x^γ, y^γ) is increasing, the inequalities must be strict for a positive measure of (s_1, s_2) .

A symmetric argument for team 2 and the fact that the game is zero sum give

$$\phi_\gamma(s_1, s_2) \leq \bar{\phi}_0(s_1, s_2), \forall (s_1, s_2),$$

again with the inequalities being strict for a positive measure of (s_1, s_2) . It thus follows that

$$\mathbb{E}_{(s_1, s_2)} [\underline{\phi}_0(s_1, s_2)] < \mathbb{E}_{(s_1, s_2)} [\phi_\gamma(s_1, s_2)] < \mathbb{E}_{(s_1, s_2)} [\bar{\phi}_0(s_1, s_2)].$$

Since the γ -auction is non-arbitrary, the proof is complete. \square

What about the divide-and-choose rule? For this rule to be *ex ante* symmetric, the divider and the chooser must be selected by a coin flip. Hence, even though a pure strategy equilibrium may exist for a realized coin flip, the probability of winning at the start of the overtime need not be non-random. For this reason, in general the divide-and-choose rule is not non-arbitrary. In fact, we can present a case favoring an auction over the divide-and-choose rule.

Proposition 3. *In a divide-and-choose rule, the chooser is better off on average than the divider. A γ -auction is fairer than a divide-and-choose rule (with a coin flip).*

Proof. Consider the divide-and-choose rule. Suppose a coin flip picks team 1 to be the divider. Suppose in the ensuing equilibrium, the yard line proposed by team 1 is distributed over support X_d according to $K(x)$. For each $x \in X_d$, let $L(s_1|x)$ be the cdf of team 1's type proposing x . (Hence, $F_1(s_1) = \int_x dL(s_1|x)dK(x)$). Let $\phi_d^-(s_1, s_2)$ denote the team 1's equilibrium payoff (i.e., winning chance) in state (s_1, s_2) when it is the divider. Likewise, we use $\phi_d^+(s_1, s_2)$ to

denote team 1's equilibrium payoff in state (s_1, s_2) when it is the chooser.

In equilibrium, observing $x \in X_d$ must lead team 2 with any type say s_2 to infer correctly about the conditional distribution of team 1's type making that proposal, and chooses a better option. Consequently, team 1's equilibrium payoff condition on proposing x is

$$\min \left\{ \int_{s_1} (a(x) + \alpha(s_1, s_2)) dL(s_1|x), \int_{s_1} (b(x) + \beta(s_1, s_2)) dL(s_1|x) \right\}.$$

Hence, for that s_2 ,

$$\begin{aligned} \mathbb{E}_{s_1} [\phi_d^-(s_1, s_2)] &= \int_{x \in X_d} \min \left\{ \int_{s_1} (a(x) + \alpha(s_1, s_2)) dL(s_1|x), \int_{s_1} (b(x) + \beta(s_1, s_2)) dL(s_1|x) \right\} dK(x). \end{aligned} \tag{1}$$

Now consider γ -auction characterized by a pure strategy equilibrium $(x^\gamma(\cdot), y^\gamma(\cdot))$. In that auction equilibrium, team 1 has an option of bidding precisely the yard line it would have proposed in an equilibrium of the divide-and-choose rule when it is the divider. Hence, for each s_2 ,

$$\begin{aligned} \mathbb{E}_{s_1} [\phi_\gamma(s_1, s_2)] &\geq \int_{x \in X_c} \left[\int_{s_1} (a(\gamma x + (1 - \gamma)y^\gamma(s_2)) + \alpha(s_1, s_2)) dL(s_1|x) \cdot 1_{\{x \leq y^\gamma(s_2)\}} \right. \\ &\quad \left. + \int_{s_1} (b(\gamma y^\gamma(s_2) + (1 - \gamma)x) + \beta(s_1, s_2)) dL(s_1|x) \cdot 1_{\{x > y^\gamma(s_2)\}} \right] dK(x) \\ &\geq \int_{x \in X_c} \min \left\{ \int_{s_1} (a(x) + \alpha(s_1, s_2)) dL(s_1|x), \int_{s_1} (b(x) + \beta(s_1, s_2)) dL(s_1|x) \right\} dK(x) \\ &= \mathbb{E}_{s_1} [\phi_c^-(s_1, s_2)] \geq \mathbb{E}_{s_1} [\underline{\phi}_c(s_1, s_2)], \end{aligned}$$

where the first inequality follows from the fact that mimicking the strategy of the divider is not necessarily an equilibrium strategy in the auction game, the second follows from the fact that $a(\cdot)$ is increasing and $b(\cdot)$ is decreasing, and the last equality follows from (1). Again given the monotonicity of (x^γ, y^γ) , the inequalities must be strict for a positive measure of s_2 .

It follows that

$$\mathbb{E}_{(s_1, s_2)} [\phi_\gamma(s_1, s_2)] > \mathbb{E}_{(s_1, s_2)} [\phi_c^-(s_1, s_2)] \geq \mathbb{E}_{(s_1, s_2)} [\underline{\phi}_c(s_1, s_2)].$$

A symmetric argument for team 2 when it is a divider and the fact that the game is zero sum give

$$\mathbb{E}_{(s_1, s_2)} [\phi_\gamma(s_1, s_2)] < \mathbb{E}_{(s_1, s_2)} [\phi_c^+(s_1, s_2)] \leq \mathbb{E}_{(s_1, s_2)} [\bar{\phi}_c(s_1, s_2)].$$

The string of inequalities proves both statements (recalling that a γ -auction is non-arbitrary). \square

Although some favors the divide-and-choose rule (see Kraska (2004)), the appeal of this rule does not extend to the more realistic incomplete information setting. The divide-and-choose procedure does not entail an envy-free division in that setting since the chooser's choice will likely cause the divider to regret her division *ex post*. As Proposition 3 confirms, the rule

is unfair to the divider, for she is likely to be taken advantage of by the chooser. Coin tossing to determine the roles will restore *ex ante* fairness, but will not fix the *ex post* fairness, and thus will likely have the same unappeal as the current system, albeit in a diminished scale. As the proposition suggests, auctions appear to provide a more fairer and thus desirable outcome.

5. Conclusion

When the Wall Street Journal ran an online forum on this issue (Wall Street Journal, 2003), Tom Donahoe, then president and general manager of the Buffalo Bills, wrote “[w]e don’t like the current system. It just seems that too much depends on the coin flip - who wins it and who loses it.” Donahoe wrote of the auction proposal being “way too revolutionary for the National Football League. It sounds more like a videogame to me. I couldn’t see that happening.” However, the NFL’s steady increase of coaching decisions to enhance fairness, e.g., the challenge rule for instant replay reviews of plays, are similar in spirit.

Although the current paper has framed its analysis in the context of the NFL overtime, the implications of our analysis goes much beyond that setting. The issue of fair division arises in many real world contexts ranging from dispute resolution, resource allocation to international trade negoti-

ation. While the extant research on fair division assumes that players have full information about the objects being divided, this condition is hardly realistic in many relevant situations. The current paper shows that accounting for the more realistic common value features changes the performances of alternative division procedures that are known to produce the same (i.e., envy-free) outcome. Our analysis suggests that, with two players, auctions tend to provide an *ex post* fairer outcome than the much heralded divide-and-choose procedure in such a setting.

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